

Lecture notes on risk management, public policy, and the financial system

Assessing Value-at-Risk

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Backtesting of VaR

Critiques of VaR

Backtesting of VaR

- Overview

- Unconditional coverage test procedure

- Examples of backtesting

- Limitations of the unconditional coverage test

Critiques of VaR

Challenges in validating VaR

- How do we measure “poor performance” of VaR? → **model risk**
- VaR **backtesting**: type of **model validation**
- VaR not a point forecast, but statement about distribution of future outcomes
- VaR **exceedance**, **exception** or **excession**: event the portfolio loss exceeds the VaR
 - Loss over the VaR horizon is compared with VaR computed just prior
 - E.g. for daily VaR, compare VaR reported at close of trading with loss over subsequent trading day
- For single position, exceedance can be defined in terms of return: for each of T observations,
 - Parametric: compare realized return with estimated volatility
 - Historical simulation: compare realized log or arithmetic return with quantile of historical sample
- Practical problem: portfolio is likely to be changing over time
 - Backtest comparison assume static portfolio

Testable dimensions of VaR

Unconditional coverage: is proportion of exceedances in entire sample consistent with VaR confidence level?

Independence: frequency and timing of exceedances, e.g. absence of clustering

Magnitude of exceedances: somewhat larger or much larger than VaR?

Brief review of statistical hypothesis testing

- Formulate **statistical hypothesis** testable with available data
 - Framed as a **null hypothesis** \mathfrak{H}_0 about a distributional characteristic of the data
 - \mathfrak{H}_0 expressed through a **test statistic**, so **falsifiable** based on data
 - \mathfrak{H}_0 guides choice of test statistic; data determines its value
 - So **falsifiable** based on data
- \mathfrak{H}_0 guides choice of test statistic; data determines its value
 - So **falsifiable** based on data

Errors in statistical hypothesis testing

Type I: reject \mathfrak{H}_0 even though \mathfrak{H}_0 true

- Often referred to as “false positive”
 - Since rejection often taken as *confirmation* of a theory
 - When framed as “treatment has effect” or “factor has influence”
- **Significance level** of test: a prespecified, chosen probability of Type I error, e.g. 0.01
 - ***p*-value**: probability, if \mathfrak{H}_0 true, of having a test statistic at least as unfavorable to \mathfrak{H}_0 as that actually obtained

Type II: fail to reject \mathfrak{H}_0 even though \mathfrak{H}_0 false

- “False negative”
- **Power** of a test: probability of Type II error

Sample space of a statistical test

- **Sample space:** all the possible configurations of the data
- Identify in the sample space for a given significance level:
 - **Critical** or **rejection region** within which \mathfrak{H}_0 rejected
 - **Acceptance** or **non-rejection region** within which \mathfrak{H}_0 *not* rejectedis complement in sample space of critical region
- Sample \in critical region leads to test statistic with p -value $<$ significance level

Statistical framework for unconditional coverage test

- VaR associated with a confidence level α
- VaR model accurate \Rightarrow exceedances occur \approx every $(1 - \alpha)^{-1}$ periods
 - For example, with daily VaR at 95 percent, expect ≈ 1 per month
 - \rightarrow Null hypothesis \mathfrak{H}_0 : exceedance frequency or fraction of exceedances = $1 - \alpha$
- Backtest is a sequence of comparisons of *current* VaR estimate with P&L realized *at the VaR forecast horizon*
- Under \mathfrak{H}_0 , comparisons are Bernoulli trials/random variables:

with probability $\left\{ \begin{array}{c} 1 - \alpha \\ \alpha \end{array} \right\}$ result is $\left\{ \begin{array}{c} 1 \text{ (VaR exceedance)} \\ 0 \text{ (VaR not exceeded)} \end{array} \right\}$

- And independently and identically distributed (i.i.d.)
 - In reality, *clustered* exceedances are routine
- \mathfrak{H}_0 doesn't state returns are lognormal, just that VaR procedure accurate for confidence level α

Test statistic of unconditional coverage test

- **Likelihood function** of T i.i.d. observations of VaR forecast and subsequent realized loss:

$$L(\alpha; x) = (1 - \alpha)^x \alpha^{T-x}$$

- x is the number of exceedances out of T
- $L(\alpha)$: probability of x in-sample exceedances if exceedance probability $1 - \alpha$
- **Maximum likelihood estimator** of α is $1 - \frac{x}{T}$
 - Likelihood function then takes on value

$$L\left(\frac{x}{T}; x\right) = \left(\frac{x}{T}\right)^x \left(1 - \frac{x}{T}\right)^{T-x}$$

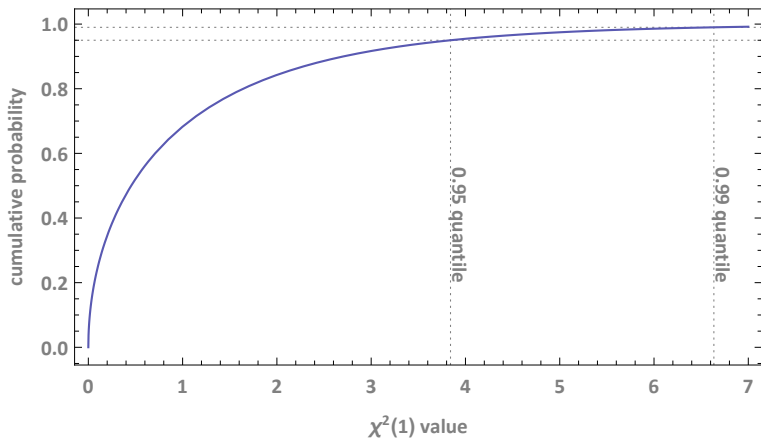
- The test statistic is the **log likelihood ratio**

$$\begin{aligned} & 2 \left\{ \ln \left[L\left(\frac{x}{T}; x\right) \right] - \ln [L(\alpha; x)] \right\} \\ & = 2 \left\{ \ln \left[\left(\frac{x}{T}\right)^x \left(1 - \frac{x}{T}\right)^{T-x} \right] - \ln [(1 - \alpha)^x \alpha^{T-x}] \right\} \end{aligned}$$

Distribution of unconditional coverage test statistic

- Test statistic measures distance between data and model prediction
 - Log of ratio of what we observe to what \mathfrak{H}_0 leads us to expect
- Follows a χ^2 **distribution** (for large enough T) if \mathfrak{H}_0 is true
 - With one **degree of freedom** (df), for the one parameter α
 - χ^2 test a standard approach to assessing **goodness of fit** of a distributional hypothesis
 - In this case, exceedances i.i.d. Bernoulli trials with parameter α
- p -value: probability, if \mathfrak{H}_0 true, of a test statistic greater than or equal to that actually obtained in the sample
 - I.e. 1 minus cumulative probability of a $\chi^2[1]$ variate with a value equal to the test statistic
- Independence requirement → non-overlapping observations if risk horizon $>$ observation frequency

$\chi^2[1]$ distribution



Cumulative distribution function of a χ^2 variate with one degree of freedom.

Significance level	0.95	0.99
Critical value	3.8415	6.6349

Critical value and acceptance range

- Reject \mathfrak{H}_0 only if test statistic $>$ critical value
 - Critical value is a quantile of $\chi^2[1]$, the χ^2 distribution with 1 df
 - Quantile is chosen to correspond to significance level of backtest
- \rightarrow **Acceptance range**: range of number of exceedances s.t. test statistic $<$ critical value
 - If number of exceedances falls *outside* acceptance range, *reject* null hypothesis
 - Too many *or* too few exceedances \rightarrow high value of test statistic
 - But caveat: χ^2 nonetheless a **one-tailed test**
- **Example**: 1 year (252 daily observations), VaR confidence level 0.99

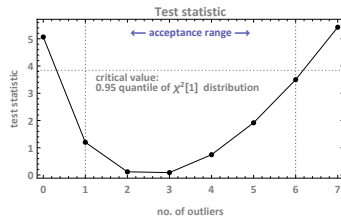
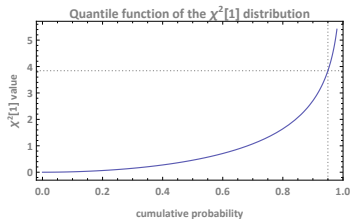
No. of exceedances	0	3	10
Test statistic	5.0654	0.0870	12.8331
χ^2 cumulative probability	0.9756	0.2320	0.9997
<i>p</i> -value	0.0244	0.7680	0.0003

- Zero exceedances results in rejection of \mathfrak{H}_0 at a significance level of 0.95, but not 0.99

Significance and confidence levels in the test

- Confidence level of backtest is distinct from confidence level of VaR
 - *Confidence* level of *VaR* enters into test statistic (together with number of observations, number of exceedances)
 - *Significance* level of *backtest* determines χ^2 quantile to compare (together with number of degrees of freedom)
- Acceptance range depends on significance level of backtest
 - Acceptance range is wider at a higher significance level
 - Greater departure from expected exceedance count required to reject null that VaR accurate
 - Any realization outside acceptance range has p -value below significance level of backtest

Test statistic and acceptance range

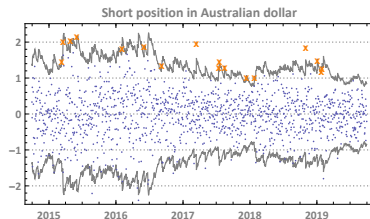
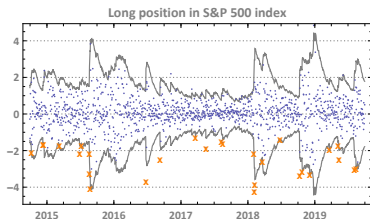


Points represent values for 1 years of daily VaR estimates; $T = 252$ and $\alpha = 0.99$ of test statistic $2 \left\{ \ln \left[\left(\frac{x}{T} \right)^x \left(1 - \frac{x}{T} \right)^{T-x} \right] - \ln \left[(1 - \alpha)^x \alpha^{T-x} \right] \right\}$ for integer values of exceedances x from 0 to 7. The acceptance range at a 95 percent confidence level is $x \in [1, 6]$.

Setting up the examples

- Unconditional coverage test of daily VaR at 99 percent confidence level
 - Using 5 years of data 30Sep2014 to 30Sep2019
 - Use parametric VaR with EWMA volatility estimate
- Assume constant position size each day, backtest in return terms
- Backtest two single-position portfolios:
 - Long position in S&P 500
 - Short position in AUD against USD
 - AUD-USD exchange rate expressed as USD price of A\$1
 - Short loss if exchange rate rises

S&P 500 and AUD-USD returns and excessions



Points denote daily returns, solid plot the 98 percent confidence level, expressed as a return and measured using a EWMA volatility estimate with a decay factor of $\lambda = 0.94$. **Orange x's** denote excessions of the VaR. Left: long position in the S&P 500 index. Right: short position in AUD against USD.

Results for the examples

- Reject \mathfrak{H}_0 for long position in S&P 500 at 0.95 and 0.99 significance levels
- Reject \mathfrak{H}_0 for short position in AUD-USD at neither 0.95 nor 0.99 significance levels

	Long S&P 500	Short AUD-USD
no. obs.	1258	1304
acceptance range (0.99 significance level)	7–20	7–20
no. excessions	28	17
% excessions	2.23	1.30
value of test statistic	14.157	1.109

Limitations of the unconditional coverage test

- Weak test: hard to reject ξ_0 unless number of observations T very large
- Disregards *size* of exceedances (→expected shortfall)
- Disregards *clustering* of exceedances (→alternative tests, return models)

Backtesting of VaR

Critiques of VaR

- Overview

- Variability of VaR estimates

- The coherence critique of VaR

Limitations of VaR

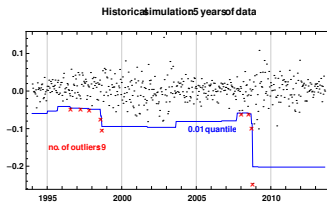
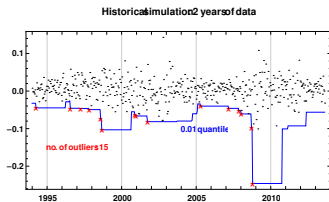
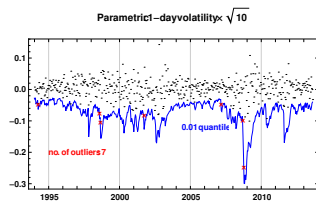
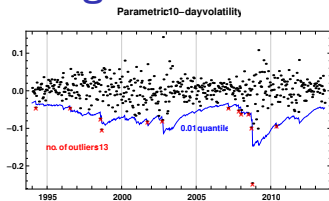
- Accuracy:
 - Inadequate treatment of frequency and size of tail risk \Rightarrow generally poor performance during crises
 - But even when no recent financial crisis, low power, i.e. hard to reject null
- VaR doesn't tell risk manager how large loss might be if VaR exceeded
 - In VaR limit system, may incentivize traders to take more risk
 - Trades may increase return, as well as probability of tail losses much larger than VaR, while increasing VaR much less
 - Can be addressed through use of (\rightarrow) **expected shortfall**
- Even if the distribution model were right: nonlinear risks, options
- The devil in the details: subtle and not-so-subtle differences in how VaR is computed \rightarrow large differences in results
- VaR is not **coherent** because it is not **subadditive**: a portfolio may have a VaR larger than the sum of the individual positions' VaR
- **Procyclicality**: widespread use of similar VaR models in setting trading limits can amplify price fluctuations

Getting whatever answer you want from VaR

- S&P 500 index Dec. 1993 to Aug. 2013
- Compute 10-day (2-week) VaR four different ways
 1. Parametric: assume log returns normally distributed
 - 1.a Using 10-day volatility, computed via exponentially weighted moving average (EWMA) using non-overlapping observations
 - 1.b Using 1-day volatility times $\sqrt{10}$
 2. Historical simulation using non-overlapping observations
 - 2.a Using 2 years of data
 - 2.b Using 5 years of data
- Express results as a return (easy to turn into a dollar amount)
- Results: large differences among approaches

<i>Technique</i>	12Mar2003	26Nov2008
Parametric: 10-day volatility	9.90	14.43
Parametric: 1-day volatility $\times \sqrt{10}$	9.03	28.75
Historical simulation: 2 years of data	8.15	24.60
Historical simulation: 5 years of data	9.66	20.15

Backtesting the four models



Backtesting VaR, 99 percent confidence level. With $T = 513$ and $\alpha = 0.99$, the acceptance range is $[2, 10]$. Points denote returns, blue plot the VaR, expressed as a return, red \times 's denote excursions.

Variability and model risk

- **Model risk:** Risk of losses due to errors in models and how applied
- Choice of VaR model can lead to over- or underestimate of risk *ex post*
- → Subject to manipulation
 - Choice of computational technique, historical lookback period
 - Distributional hypothesis, pricing models in simulations
 - Choice of risk factors, e.g. mapping resi subprime to AAA corporate
 - Mapping position and hedge to same risk factor: voil'a, no basis risk

Coherence of risk measures

- **Coherence** is a set of standards for risk measures to ensure they do not lead to perverse or counterintuitive rankings of strategies
- Defined mathematically, but implement these intuitions:
 - **Monotonicity:** if one portfolio's return is always greater than that of another, its measured risk must be smaller
 - **Homogeneity of degree one:** doubling every position in a portfolio should exactly double its measured risk
 - **Subadditivity:** the risk of a portfolio should be no greater than the sum of the risks of its constituents
 - **Translation invariance:** adding a riskless asset to a portfolio should reduce its measured risk by that same amount
- VaR doesn't satisfy the subadditivity condition

Examples of failure of subadditivity of VaR

- Examples are easy to generate: require
 - Positions susceptible to large loss, but with low probability, i.e. below $1 - \alpha$, with α the VaR confidence level
 - → Each position has zero or negative VaR
 - Positions are independent, or have low correlation, or low probability of joint event of loss
 - Loss probabilities and correlations are such that probability of *loss on at least one position* exceeds α
- **Examples** of positive-VaR portfolios at the 99 percent confidence level consisting of zero- or negative-VaR positions
 - Market-risk VaR: two option positions, short a far out-of-the-money (OTM) call and OTM put, each with probability of exercise just less than 1 percent
 - Credit-risk VaR: two loans, each with a default probability just less than 1 percent and low default correlation